Maximally Forward-Looking Core Inflation

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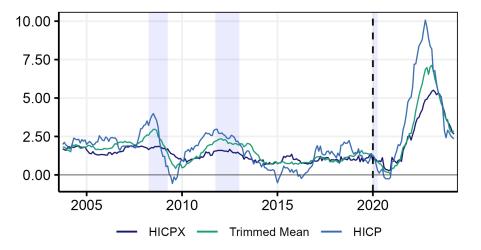
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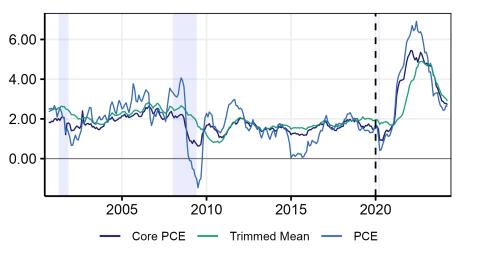
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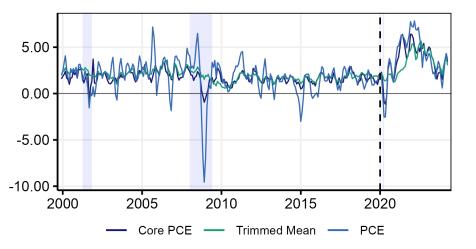
Where are we? Some measures of inflation (YoY) for the Euro Area



Where are we? Some measures of inflation (YoY) for the US



Where are we? Some measures of inflation (QoQ) for the US



Core Inflation: A Primer

How are they created?

- 1. Heuristics (permanent exclusion)
 - HICPX/PCE core (HICP/PCE minus food and energy)
 - PCE supercore (PCE core services minus housing)
- 2. Heuristics (temporary exclusion)
 - Median inflation
 - Trimmed-mean inflation
- 3. Model-based (unsupervised)
 - (US) Multivariate Core Trend from FRB-NY (Stock and Watson, 2016)
 (EA) PCCI (Bańbura and Bobeica, 2020)

What are the desirable properties?

- 1. Low bias
- 2. Low volatility
- 3. Low error in forecasting headline inflation
- 4. Highly reactive to monetary policy and labor market conditions

Contribution

- We write a simple machine learning regression problem that generates, by construction, the linear aggregation of subcomponents that is maximally predictive of future inflation conditions.
- We call it the *Assemblage Regression*.
- In the context of inflation modeling, the resulting product is Albacore, for adaptive learning-based core inflation.
- Features: (3) good forecasting performance, (1) low bias, (2) low volatility.
- Criterion (4) is the subject of another ongoing project.

Assemblage Regression

Our proposed measures are estimated via a **generalized non-negative ridge regression** where the dependent variable is future headline inflation:

Supervised Weighting with Albacore_{comps}

$$\hat{oldsymbol{w}}_{c} = rgmin_{oldsymbol{w}} \sum_{t=1}^{T-h} (\pi_{t+1:t+h} - oldsymbol{w}' oldsymbol{\Pi}_{t})^2 + \lambda ||oldsymbol{w} - oldsymbol{w}_{ ext{PCE}}||_2 \quad ext{st} \,\,oldsymbol{w} \geq 0, \,\,oldsymbol{w}' \iota = 1$$

where $\pi_{c,t}^* \equiv \hat{w}_c' \Pi_t$ and $h \in \{6, 12, 24\}$ months. Π_t are inflation subcomponents.

Supervised Trimming with Albacore_{ranks}

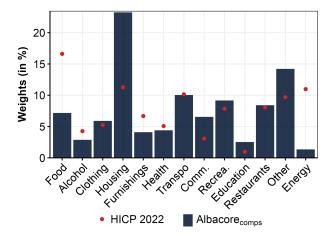
$$\hat{\boldsymbol{w}}_r = \operatorname*{argmin}_{\boldsymbol{w}} \sum_{t=1}^{T-h} (\pi_{t+1:t+h} - \boldsymbol{w}' \boldsymbol{O}_t)^2 + \lambda ||D\boldsymbol{w}||_2 \quad \mathrm{st} \ \boldsymbol{w} \geq 0$$

where $\pi_{r,t}^* \equiv \hat{w}_r' O_t$ and $h \in \{6, 12, 24\}$ months. O_t stores inflation subcomponents sorted at each *t*.

Experiments

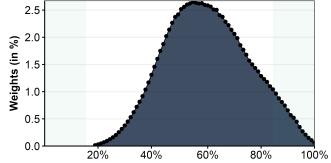
- Monthly data, predicting out-of-sample the average path of headline inflation for the next *h* ∈ {6, 12, 24} months.
- $\Pi_t \equiv \text{PCE/HICP/CPI}$ components at 3 levels of disaggregation. US: level $\in \{2, 3, 6\}$ implying $K \in \{15, 50, 215\}$ EA: level $\in \{2, 3, 4\}$ implying $K \in \{12, 39, 92\}$
- Components are expressed in QoQ (the change vs 3 months ago).
 → Shorter averages are too noisy, longer ones will not be timely in time of need.
- Benchmarks: forecast combination of headline and a wide range of existing core inflation measures
- Two test periods, with a 20 years rolling window estimation
 - 1. Pre-Covid: 2010-2019
 - 2. Post-Covid: 2020-2023
- \Rightarrow Albacore_{comps} and Albacore_{ranks} show good forecasting performance in all levels of disaggregation.

Albacore_{comps}



- High weight on *Housing* (rents, housing services), *Communication*, and *Other Services* → rather persistent components.
- *Energy* is ~excluded; *Food* << HICP weight.

Albacore_{ranks}



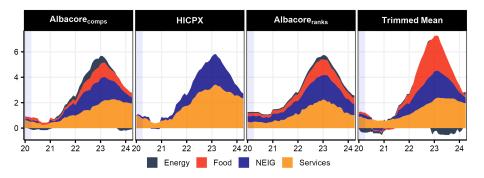
- Tails are trimmed, but we obtain highly *asymmetric* trimming.
- What goes up is more likely here to stay. Rockets and feathers? Market power? Downward rigidities?
- Monthly price distribution is asymmetric (Carroll and Verbrugge, 2019).
- The pandemic caused a shift from slightly negative to positive skeweness of distribution of price changes. Median and symmetric trimming understated (early) trends (Rich et al., 2022). Albacore_{ranks} does not suffer from this flaw.

Narratives Comparison YoY



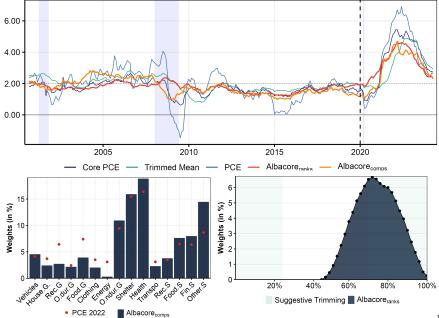
- Albacore_{ranks} shows upward pressures on inflation as early as mid-2020
- Albacore_{comps} is more in line with traditional core measures for the initial surge, but captures the turning point earlier.

Albacore Decomposition EA

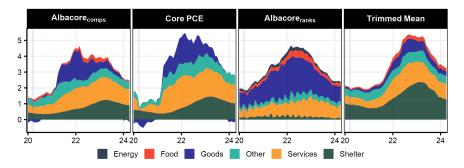


- Commodity prices (food and energy) accounted for the majority of the post-pandemic inflation acceleration, excluded a priori by HICPX.
- Albacore shows positive contributions of food components, which aligns with, e.g., Peersman (2022), highlighting the importance of global food commodity prices for euro area inflation dynamics.
- Persistence stems from services, in particular, for Albacore_{comps}, rents, restaurants and hotel services as well as other services.

Albacore for the US



Albacore Decomposition US



- Albacore_{ranks} already spots signs of upward tendencies in mid-2020 that build up as more and more sectors face negative impacts from the mix of supply chain disruptions and sustained consumer spending.
- We find high contribution of goods inflation for Albacore_{ranks} whereas the trimmed mean is driven by shelter (which is known to be lagging).
- Contrary to PCE core, neither Albacore_{comps} nor Albacore_{ranks} exclude food or energy prices a priori, and yet, they both show little contribution of the corresponding components.

Parting Words

- We devised 2 simple maximally forward-looking core inflation series.
- This is part of a broader research agenda where macro aggregates are (partly) redefined by a machine learning algorithm so to better satisfy certain (economic or statistical) evaluation criteria.
- Next step (ongoing): using Goulet Coulombe and Göbel (2023)'s MACE algorithm, optimize CPI weights so the aggregate inflation metric is maximally predictable based on the monetary policy instrument and real activity conditions.
- Final destination: Assemblage VAR

Appendix

Detour: A Primer on Ridge Regression

- A ridge regression solves what is more generally known as a penalized linear regression problem.
- The RR coefficients are obtained via

$$\hat{\boldsymbol{\beta}}_{\mathrm{Ridge}} = rgmin_{\boldsymbol{\beta}} \sum_{t=1}^{T-1} (y_{t+1} - \boldsymbol{\beta}' \boldsymbol{X}_t)^2 + \lambda ||\boldsymbol{\beta}||_2$$

where $||.||_2$ is the l_2 norm. The latter is equivalently $\sum_{k=1}^{K} \beta_k^2$ in summation notation, where K = ncol(X).

- The penalty term provides regularization by bringing in the mix the a priori that each coefficient should contribute to the fit, but modestly. In other words, it is *shrinking* coefficients towards 0.
- With a suitable $\lambda > 0$, RR curbs overfitting that would plague OLS (i.e., $\lambda = 0$), especially when *K* is large relative to *N*.
- λ is typically tuned via cross-validation, which is, in essence, a pseudo-out-of-sample evaluation metric.

Assemblage Regression Supervised Weighting with Albacore_{comps}

- Our proposed measure is estimated via a **non-negative ridge regression** where the dependent variable is future headline inflation and the regressors are inflation subcomponents.
- The permanent exclusion version, or supervised weighting of basket components, is obtained via

$$\hat{w}_{\scriptscriptstyle \mathcal{C}} = rgmin_w \sum_{t=1}^{T-h} (\pi_{t+1:t+h} - w' \mathbf{\Pi}_t)^2 + \lambda ||w - w_{\scriptscriptstyle \mathrm{PCE}}||_2 \quad \mathrm{st} \ w \geq 0, \ w' \iota = 1$$

where $\pi_{c,t}^* \equiv \hat{w}_c' \Pi_t$ and $h \in \{6, 12, 24\}$ months.

- Note the absence of an intercept which forces (i) the random walk hypothesis and (ii) $\pi_{c,t}^*$ to have the same unconditional mean as $\pi_{t+1:t+h}$ through $w_{\text{intercept}} = \bar{\pi}_{t+1:t+h} \bar{\pi}_t^* = \bar{\pi}_t \bar{\pi}_t^* \approx 0$, therefore making it already denoted in headline inflation units.
- \rightarrow Some kind of multivariate-to-univariate Hamilton filter (Hamilton, 2018).

Assemblage Regression

Supervised Trimming with Albacore_{ranks}

- This is effectively just sorting the components at each *t* and inputting them in the matrix *O*_{*t*}.
- The switch to "rank statistics space" can be formalized as $O_t = A_t \Pi_t$ where A_t is an allocation matrix.
- Rather than learning which subcomponents to include, the problem will now be learning which ranks (and with which weight) to include or exclude. Thus, we run

$$\hat{oldsymbol{w}}_r = rgmin_{oldsymbol{w}} \sum_{t=1}^{T-h} (\pi_{t+1:t+h} - oldsymbol{w}'oldsymbol{O}_t)^2 + \lambda ||Doldsymbol{w}||_2 \quad ext{st} \ oldsymbol{w} \geq 0$$

where $\pi_{r,t}^* \equiv \hat{w}_r' O_t$ and $h \in \{6, 12, 24\}$ months.

- Albacore_{ranks} will favor a smooth and adaptive trim, in contrast to
 - FRB Cleveland trimmed-mean CPI, which sets to 0 the first and last 16% of ranks, and then a weighted average of the center band is reported.
 - FRB Dallas trimmed-mean PCE inflation, which trims 24% from lower tail and 31% from the upper tail.
 - Median CPI, which keeps only one rank, the middle one.

Euro Area Results

	20	10m1 - 2019	m12	2020m1 -2023m12					
	h = 6	h = 12	h = 24	h = 6	h = 12	h = 24			
Level 2 ($K = 12$)									
Albacore _{comps}	0.94	0.84	0.75	0.96	0.99	1.02			
Albacore _{ranks}	0.90	0.84	0.78	0.95	0.94	0.95			
Level 3 ($K = 39$)									
Albacore _{comps}	0.93	0.88	0.81	1.01	1.00	1.07			
Albacore _{ranks}	0.90	0.85	0.85	0.97	0.95	0.94			
Level 4 ($K = 92$)									
Albacore _{comps}	0.96	1.00	0.85	1.06	0.97	1.04			
Albacore _{ranks}	0.92	0.87	0.87	0.93	0.93	0.91			
Benchmarks									
X_t^{bm} , $(w_0=0)$	0.95	0.90	0.78	0.98	1.02	1.09			
$X_t^{\text{bm+}}$	0.92	0.98	1.00	0.93	0.97	1.00			
$X_t^{\mathrm{bm}+}$, ($w_0=0)$	0.91	0.86	0.80	1.02	0.98	1.05			

Notes: Numbers in **bold** indicate the best model for each *level* and each *horizon* in each of the out-of-sample periods. Numbers highlighted in green show the best model per *horizon* and out-of-sample period *across levels*.

US Results

	20	10m1 - 2019ı	m12	2020m1 -2023m12					
i	h = 6	h = 12	h = 24	h = 6	h = 12	h = 24			
Level 2 ($K = 15$)									
Albacore _{comps}	1.08	1.13	1.12	0.70	0.63	0.88			
Albacore _{ranks}	0.99	0.98	0.87	0.57	0.59	0.69			
Level 3 ($K = 50$)									
Albacore _{comps}	1.10	1.14	1.06	0.80	0.75	1.01			
Albacore _{ranks}	0.93	0.87	0.73	0.61	0.56	0.63			
Level 6 ($K = 215$)									
Albacore _{comps}	1.10	1.19	1.16	0.84	0.71	0.99			
Albacore _{ranks}	0.88	0.84	0.77	0.70	0.62	0.67			
Benchmarks									
$X_t^{\text{bm}}, (w_0 = 0)$	1.03	1.03	0.96	0.79	0.80	0.93			
$X_t^{\text{bm+}}$	1.04	1.04	1.02	1.18	1.14	1.09			
$X_t^{\mathrm{bm}+}$, ($w_0 = 0$)	1.07	1.07	0.97	0.86	0.87	0.96			

Notes: Numbers in **bold** indicate the best model for each *level* and each *horizon* in each of the out-of-sample periods. Numbers highlighted in green show the best model per *horizon* and out-of-sample period *across levels*.

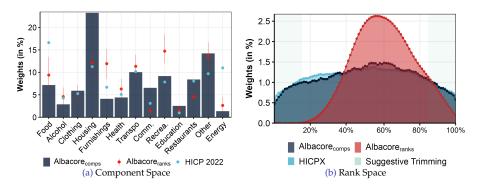
From One Space to Another A time-varying parameters view

• The supervised trimming regression can be written in components space as

$$\hat{w}_r = \operatorname*{argmin}_{w} \sum_{t=1}^{T} (\pi_{t+1:t+h} - \underbrace{w_r' A_t}_{w_{c,t}} \mathbf{\Pi}_t)^2 + \lambda ||w_r||^2 \quad \text{st } w_r \ge 0$$

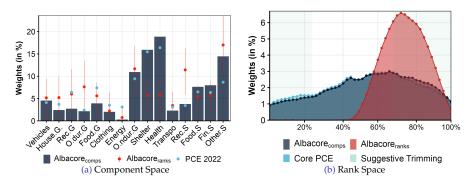
- Thus, fitting the regression in rank space implies *time-varying parameters* in the components space.
- Conversely, the regression in the original components space implies time-varying trimming via $w_{r,t} = w'_c A_t^{-1}$
- Traveling from one space to the other could be informative

From One Space to Another A look at our main EA models



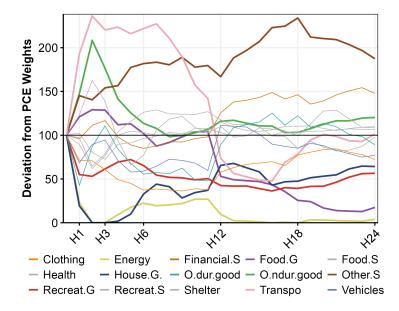
- The asymmetric trimming of Albacore_{ranks} leads to low weights for components in the left tail.
- Albacore_{comps} covers the whole distribution; similar to HICPX, however, the latter being reweighted with actual HICP weights.

From One Space to Another A look at our main US models



- The asymmetric trimming of Albacore_{ranks} leads to low weights for components in the left tail (which are mostly Energy, Food, Clothing).
- Albacore_{comps} covers the whole distribution but left-skewed; similar to Core PCE, however, the latter being reweighted with actual PCE weights.

Weights Curve Across Horizons: US Albacore_{comps}



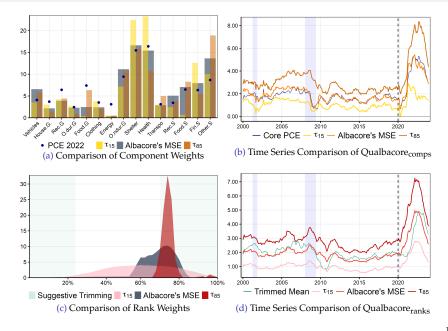
Qualbacore US Results

	2010m1-2019m12					2020m1-2023m12							
	h = 3				h = 6			h = 3			h = 6		
	τ_{15}	MSE	τ_{85}	τ_{15}	MSE	$ au_{85}$	τ_{15}	MSE	τ_{85}	τ_{15}	MSE	τ_{85}	
Qualbacore _{comps} Qualbacore _{ranks}	0.97 0.94	1.06 0.92	1.29 0.95	1.10 0.94		1.21 0.86	1.24 0.99	1.07 0.80	0.75 0.36	1.23 0.97	0.97 0.72	0.71 0.40	
Benchmarks													
$egin{aligned} & X_t^{\mathrm{bm}}, & (w_0=0) \ & X_t^{\mathrm{bm+}} \ & X_t^{\mathrm{bm+}}, & (w_0=0) \end{aligned}$	0.96 1.03 1.02	0.99 1.03 1.02	1.12 1.13 1.13	0.97 1.07 1.07	1.01 1.03 1.04	1.11 1.11 1.13	1.04 1.14 1.14	0.96 1.22 1.02	0.81 0.81 0.81	1.00 1.01 1.01	0.94 1.07 0.95	0.90 0.83 0.90	

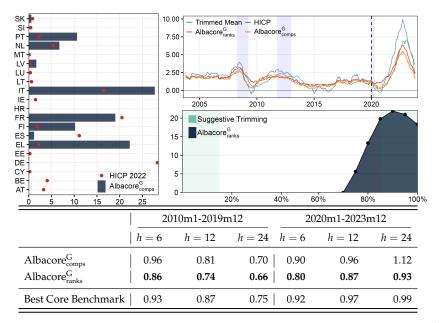
Table: Quantile Forecasting Performance of Albacore for the US

Notes: The table presents root mean square error (RMSE) relative to the AR model. Our benchmarks are: $X_t^{\text{bm}} = [\text{PCE}_t \text{ PCEcrore}_t \text{ PCEtrim}_t]$ without an intercept (i.e., $w_0 = 0$), $X_t^{\text{bm}+}$ with and without an intercept. Numbers in **bold** indicate the best performing model for each pair of horizon and loss function. Numbers in **green** highlights the loss function for which there is the largest improvement with respect to the benchmark, if applicable. In the results Section, we present level 6 for Qualbacore_{comps} (with K = 215) and level 3 for Qualbacore_{ranks} (with K = 50). Note that unlike previous results, we consider an expanding window from 1990 to increase the inevitably scarce number of observations in the tails.

Quantile regression extension for the US (h = 6)



Geographic Assemblage



Properties of Underlying Inflation Measures for US

Inflation series	Bias Full Pre-cov			Volatility Full Pre-cov		ient of var. Pre-cov		Lead/lag corr. Full Pre-cov		
	1 un	110-000	1 un	110-007	Full	110-007	i un	110-000		
3Mo3M										
Core PCE	-0.12	-0.09	0.57	0.36	0.59	0.39	0	0		
Trimmed Mean (FedDallas)	0.02	0.14	0.42	0.29	0.41	0.27	-8	0		
Core excl. housing	2.35	2.37	5.45	1.83	2.52	0.80	1	2		
Median CPI (FedČleveland)	0.59	0.59	0.61	0.38	0.46	0.29	-8	-9		
Trimmed Mean (FedCleveland)	0.33	0.31	0.63	0.40	0.53	0.34	0	0		
Sticky Core (FedAtlanta)	0.51	0.56	0.57	0.37	0.45	0.28	-9	-11		
Albacoreranks	-0.03	0.03	0.40	0.22	0.40	0.21	0	0		
Albacorecomps	-0.08	0.00	0.48	0.34	0.49	0.34	0	0		
ΥοΥ										
Core PCE	-0.13	-0.11	0.69	0.39	0.49	0.22	0	0		
Trimmed Mean (FedDallas)	0.01	0.13	0.56	0.46	0.37	0.23	-7	-9		
Core excl. housing	2.42	2.43	2.93	2.03	0.91	0.46	2	3		
Median CPI (FedČleveland)	0.55	0.56	0.82	0.61	0.43	0.25	-8	-11		
Trimmed Mean (FedCleveland)	0.30	0.29	0.82	0.55	0.47	0.25	-4	-6		
Sticky Core (Fed Atlanta)	0.48	0.53	0.74	0.59	0.40	0.24	-8	-12		
Albacoreranks	-0.04	0.01	0.54	0.28	0.36	0.15	0	-1		
Albacorecomps	-0.09	0.00	0.60	0.50	0.41	0.26	0	2		

Notes: Cells shaded in green indicate the three best performing models for each criterion. The evaluation is based on the sample ranging from 2000ml to 2023m12 for the full sample and from 2000ml to 2019m12 for the pre-Covid case. Bias is determined as the difference between the long-run average of the respective series and PCE headline. Volatility refers to the standard deviation of each series relative to the standard deviation of headline. Coefficient of variation is the ratio between standard deviation and long-run average of each measure. Positive (negative) numbers for the lead/lag correlation refer to leads (lags) with the highest corse-correlation between each measure and PCE headline.

Properties of Underlying Inflation Measures for EA

Inflation series	Bias			latility		ient of var.		Lead/lag corr.		
	Full	Pre-cov	Full	Pre-cov	Full	Pre-cov	Full	Pre-cov		
PoP										
HICPX	-0.46	-0.29	0.53	0.37	0.72	0.36	-6	0		
Trimmed Mean 30%	-0.22	-0.06	0.68	0.60	0.81	0.51	0	0		
HICP excl. energy	-0.19	-0.11	0.67	0.50	0.78	0.44	-6	0		
HICP excl energy & unpr food	-0.25	-0.15	0.64	0.45	0.77	0.40	-6	0		
Supercore	-0.25	-0.09	0.55	0.37	0.68	0.33	-6	-2		
PCCI excl. energy	5.28	4.89	1.43	1.02	0.43	0.21	-5	-1		
PCCI	5.91	5.43	1.68	1.27	0.47	0.24	-1	0		
Albacoreranks	-0.17	-0.01	0.48	0.29	0.59	0.25	-2	-1		
Albacorecomps	-0.23	-0.06	0.52	0.43	0.65	0.38	0	0		
ΥσΥ										
HICPX	-0.48	-0.29	0.55	0.46	0.63	0.31	-5	-6		
Trimmed Mean 30%	-0.24	-0.07	0.72	0.70	0.72	0.41	-3	-3		
HICP excl. energy	-0.21	-0.13	0.73	0.59	0.73	0.36	-4	-2 -2 -5		
HICP excl energy & unpr food	-0.26	-0.15	0.70	0.56	0.71	0.34	-4	-2		
Supercore	-0.27	-0.11	0.61	0.46	0.64	0.29	-4			
PCCI excl. energy	-0.25	-0.04	0.44	0.36	0.44	0.21	0	0		
PCCI	-0.09	0.10	0.51	0.46	0.48	0.24	2	2		
Albacoreranks	-0.18	-0.01	0.55	0.36	0.55	0.22	-3	-4		
Albacorecomps	-0.25	-0.06	0.57	0.53	0.60	0.33	-2	1		

Notes: Cells shaded in green indicate the three best performing models for each criterion. The evaluation is based on the sample ranging from 2000m1 to 2019m12 for the pre-Covid case. Bias is determined as the difference between the long-run average of the respective series and HICP. Volatility refers to the standard deviation of each series relative to the standard deviation of heads. Coefficient of variation is the ratio between standard deviation and long-run average of each measure. Positive (negative) numbers for the lead/lag correlation refer to leads (lags) with the highest cross-correlation between each measure and HICP.

Canada Results

	20	10m1 - 2019ı	m12	2020m1 - 2023m12					
<i>i</i>	h = 6	h = 12	h = 24	h = 6	h = 12	h = 24			
Level 3 ($K = 19$)									
Albacore _{comps}	1.05	1.13	1.37	0.68	0.63	0.85			
Albacore _{ranks}	1.05	1.13	1.15	0.74	0.74	0.78			
Level 4 ($K = 49$)									
Albacore _{comps}	1.12	1.23	1.24	0.54	0.47	0.72			
Albacore _{ranks}	1.07	1.15	1.18	0.68	0.67	0.76			
Level 5 ($K = 87$)									
Albacore _{comps}	1.10	1.22	1.24	0.69	0.55	0.71			
Albacore _{ranks}	1.05	1.12	1.27	0.76	0.71	0.74			
Benchmarks				[
X_t^{bm} , $(w_0=0)$	1.10	1.34	1.51	0.72	0.78	0.86			
$X_t^{\mathrm{bm}+}$	1.00	1.00	1.00	1.00	1.00	1.00			
$X_t^{\mathrm{bm}+}$, $(w_0=0)$	1.08	1.24	1.35	0.74	0.80	0.85			

Notes: Numbers in **bold** indicate the best model for each *level* and each *horizon* in each of the out-of-sample periods. Numbers highlighted in green show the best model per *horizon* and out-of-sample period *across levels*.

Canada

